

# Discrete Causal Representations from Heterogeneous Domains

## A Bayesian Approach with Social Survey Applications

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**Causal representation learning** aims to extending causal modeling to settings in which the causal variables of interest are latent and only indirectly observed. Existing work in this area has primarily focused on identifiability under idealized assumptions, while practical methods for finite, real-world data remain less developed. **This work is a step towards bridging that gap and provides a successful application of a complete unsupervised CRL pipeline to complex real-world data.**

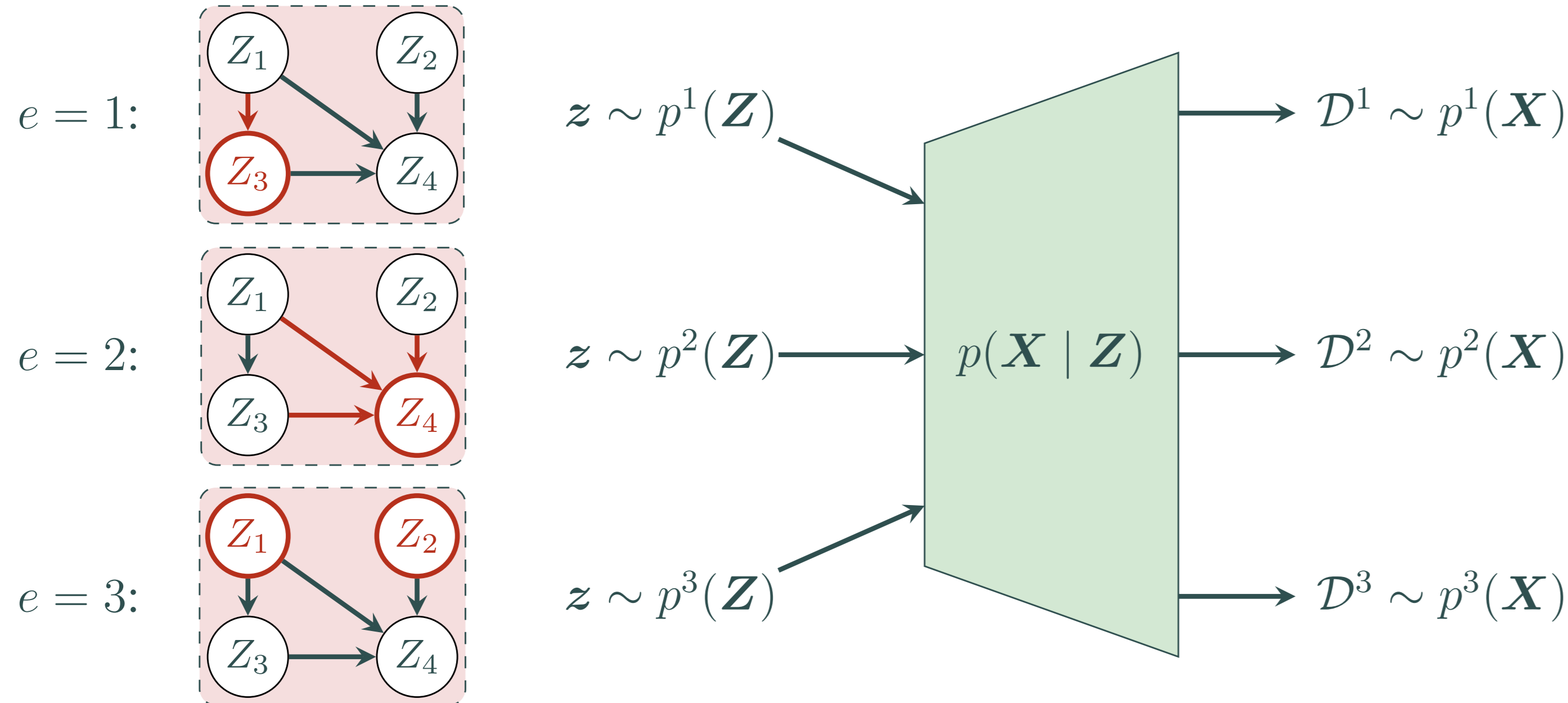


Figure 1: Multi-domain CRL with soft, multi-node interventions

### Problem Setting

- We **observe samples**  $\mathbf{x}^{e,j} \in \mathbb{R}^D$  from environments  $e \in [E]$
- Each  $\mathbf{x}^{e,j}$  is generated by some **discrete latent variables**  $z^{e,j} \in [K]^L$
- Changes across environments are caused by **sparse interventions**.

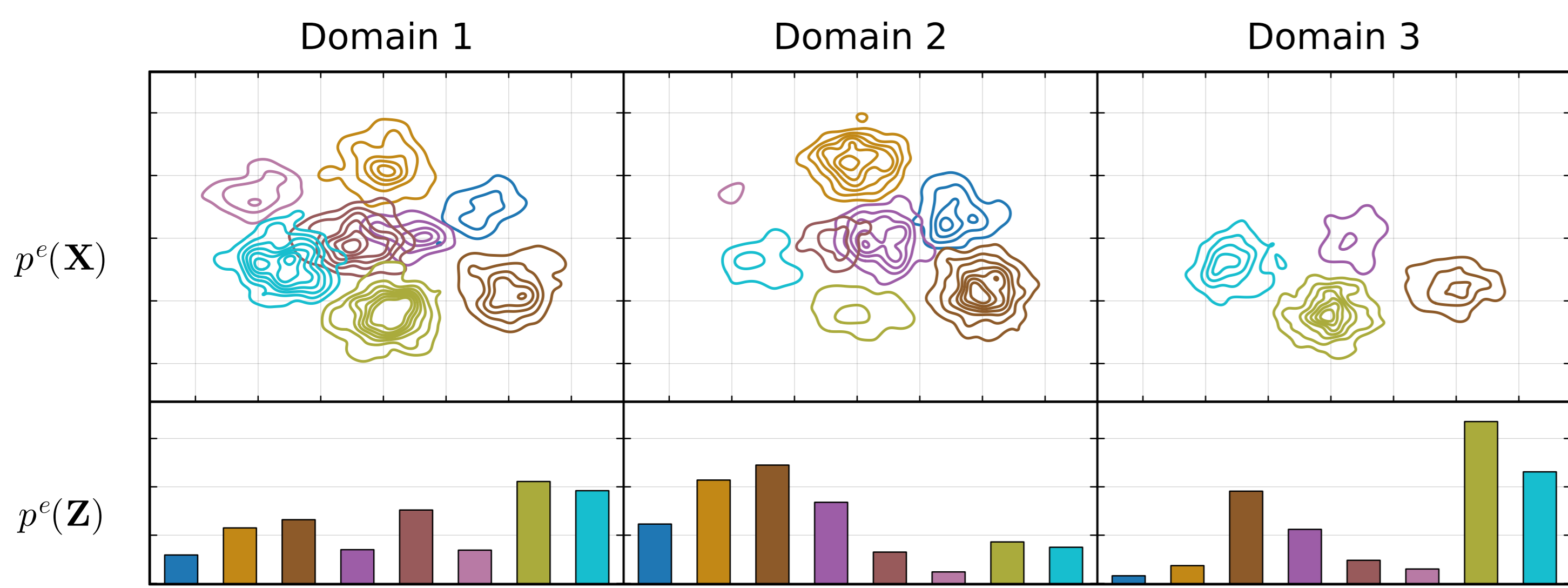


Figure 2: Observations and latents in multi-domain data

### Measurement Model

$$\mathbf{x}^{e,j} \sim \mathcal{N}_D(\mathbf{m}_0 + \sum_{\ell \in [L]} \mathbf{m}_\ell z_\ell^{e,j}, \text{diag}(\sigma^2))$$

$$\mathbf{m}_\ell \sim \mathcal{N}_D(\mathbf{0}, \mathbf{I}), \quad \sigma_d^2 \sim \text{Inv-}\Gamma(\alpha_\sigma, \beta_\sigma)$$

### Latent Mechanisms

$$\theta_\ell(\mathbf{p}\mathbf{a}_\ell) \sim \mathcal{N}_K(\mathbf{0}, \Sigma_\theta)$$

$$\mathcal{I}_\ell^e \sim \text{Beta-Bernoulli}(\alpha_c, \beta_c)$$

$$\delta_\ell^e(\mathbf{p}\mathbf{a}_\ell) \sim p(\delta | \Sigma_\Delta, \varepsilon) \propto \mathcal{N}_K(\mathbf{0}, \Sigma_\Delta) \cdot \mathbb{1}\left\{\frac{1}{2}\delta^\top \mathbf{H}\delta \geq \varepsilon\right\}$$

$$z_\ell^{e,j} \sim \text{Categorical}\left(\text{softmax}\left(\theta_\ell(\mathbf{p}\mathbf{a}_\ell^{e,j}) + \mathcal{I}_\ell^e \delta_\ell^e(\mathbf{p}\mathbf{a}_\ell^{e,j})\right)\right)$$

### Sparse Detectable Interventions

- **Sparsity:** The Beta-Bernoulli prior on  $\mathcal{I}_\ell^e$  induces sparse interventions.
- **Detectability:** The truncation prior on  $\delta_\ell^e$  rules out arbitrarily small shifts by requiring every active intervention to induce an approximate KL change of at least  $\varepsilon$ :  $\text{KL}(\text{softmax}(\theta) \| \text{softmax}(\theta + \delta)) \approx \frac{1}{2}\delta^\top \mathbf{H}\delta \geq \varepsilon$ .

### Posterior Landscape

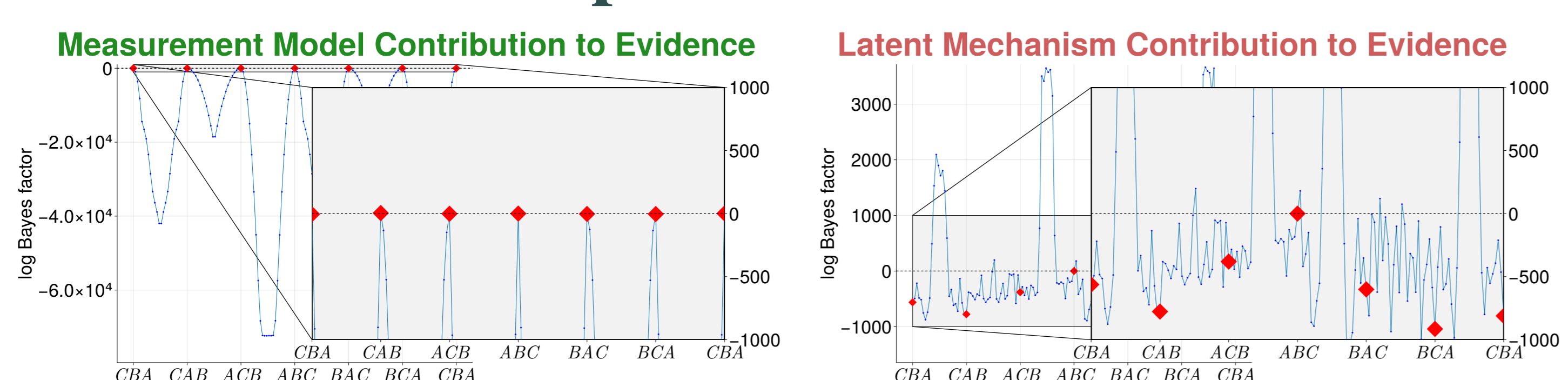


Figure 3: Multimodality across permutations of the causal ordering

- Competing posterior modes correspond to different causal orderings.
- The **Measurement Model** is permutation-invariant, creating the modes.
- The **Latent mechanisms** identifies the true causal order among them.

### Case Study I: World Values Survey (WVS)

- We analyze wavs 7 of the WVS<sup>2</sup>, a large-scale cross-national survey.
- We recover interpretable structure that matches global cultural variation.

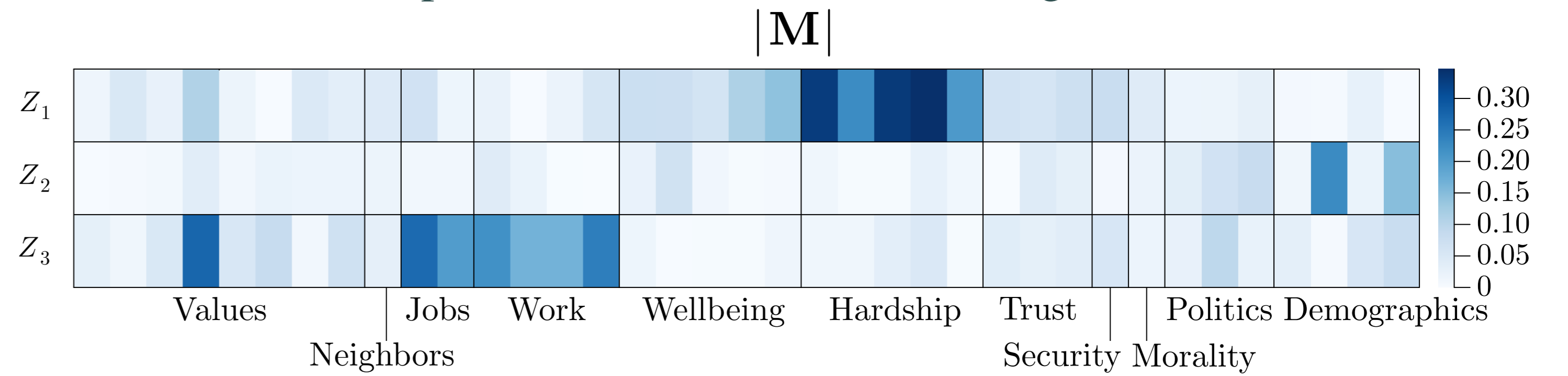


Figure 4: Measurement matrix from latent variables to survey responses.

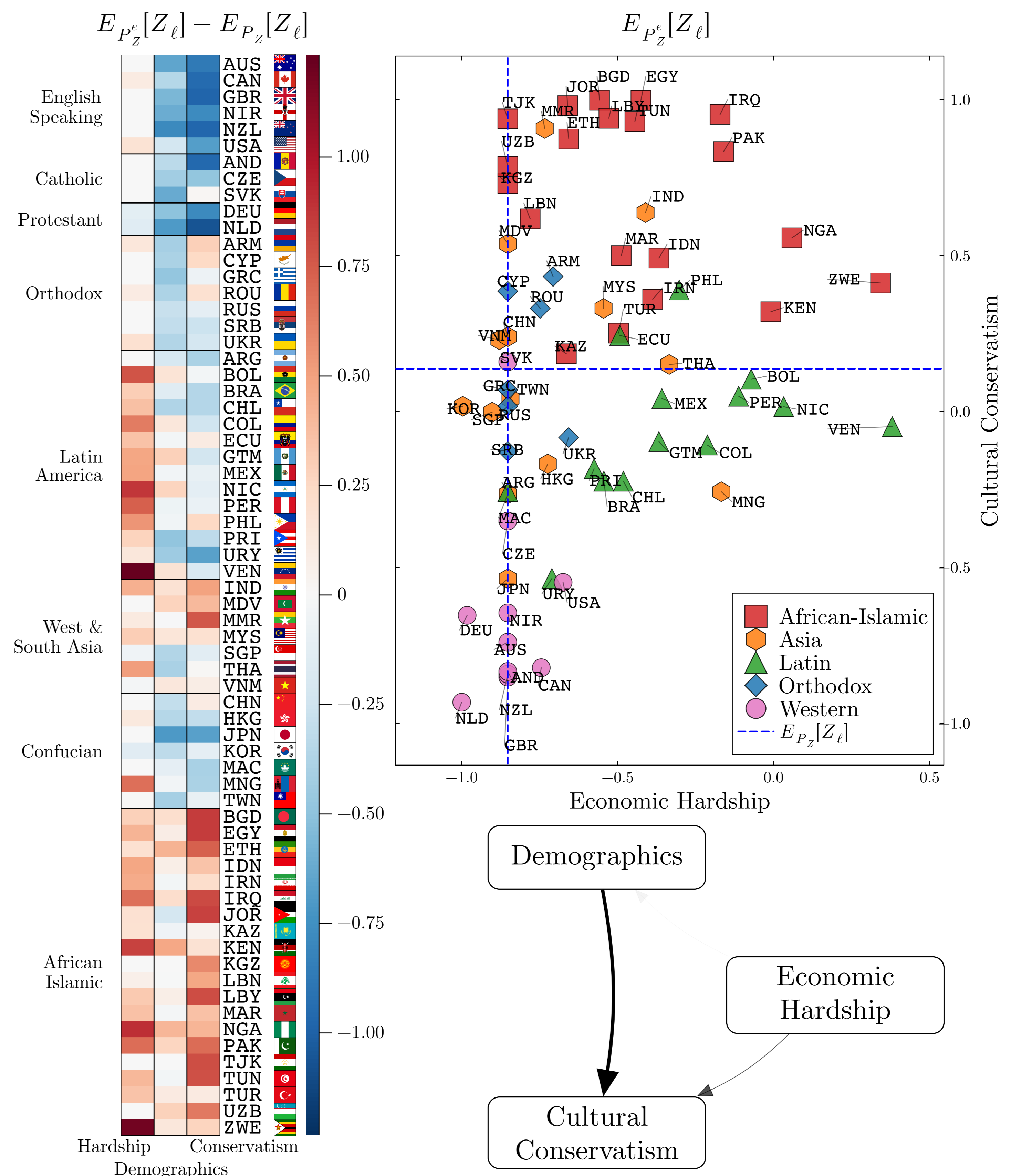


Figure 5: Marginal distributions and causal influence of latent concepts.

### Sequential Monte Carlo Sampling (SMCS)

- The model admits efficient **closed-form conjugate updates**.
- Standard MCMC chains can get trapped in one causal ordering mode.
- SMCS<sup>3</sup> uses many weighted particles to explore modes in parallel.
- Likelihood tempering allows latent mechanisms to guide inference.

$$p_{T_k}(\underbrace{\Theta, \mathcal{I}, \Delta}_{\Phi}, \underbrace{M, \sigma^2}_{\Psi}, \mathbf{Z}^\varepsilon | \mathbf{X}^\varepsilon) \propto p(\Phi)p(\Psi)p(\mathbf{Z}^\varepsilon, \mathbf{X}^\varepsilon | \Phi, \Psi)^{\frac{1}{T_k}}$$

#### Algorithm 1: Sequential Monte Carlo Sampler with Likelihood Tempering

- Require:** Temperatures  $\{\infty = T_0 > \dots > T_K = 1\}$ , number of particles  $S$ , tempered posterior  $p_{T_k}(\Phi, \Psi, \mathbf{Z} | \mathbf{X})$ , and tempered  $p_{T_k}$ -invariant MCMC kernels (GIBBSUPDATE).
- 1:  $(\Phi, \Psi, \mathbf{Z})_0^{(s)} \stackrel{\text{i.i.d.}}{\sim} p_{T_0}(\Phi, \Psi, \mathbf{Z})$  for all  $s \in [S]$  ▷ Initialize particles
  - 2:  $w_0^{(s)} \leftarrow 1/S$  for all  $s \in [S]$  ▷ Initialize weights
  - 3: **for**  $k = 1$  **to**  $K$  **do**
  - 4: Set weights  $w_k^{(s)} \propto \frac{p_{T_k}(\Phi^{(s)}, \Psi^{(s)}, \mathbf{Z}^{(s)} | \mathbf{X})}{p_{T_{k-1}}(\Phi^{(s)}, \Psi^{(s)}, \mathbf{Z}^{(s)} | \mathbf{X})}$  for all  $s \in [S]$  ▷ Reweight
  - 5:  $\{(\Phi, \Psi, \mathbf{Z})_k^{(s)}, w_k^{(s)}\}_{s=1}^S \leftarrow \text{RESAMPLE}\left(\{(\Phi, \Psi, \mathbf{Z})_{k-1}^{(s)}, w_{k-1}^{(s)}\}_{s=1}^S\right)$  ▷ Resample
  - 6:  $(\Phi, \Psi, \mathbf{Z})_k^{(s)} \leftarrow \text{GIBBSUPDATE}\left((\Phi, \Psi, \mathbf{Z})_{k-1}^{(s)}, T_k\right)$  for all  $s \in [S]$  ▷ Mutate
  - 7: **end for**
  - 8: **return**  $\left\{\left((\Phi, \Psi, \mathbf{Z})_K^{(s)}, w_K^{(s)}\right)\right\}_{s=1}^S$  as approximate posterior samples

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[3] Pierre Del Moral, Arnaud Doucet, and Ajay Jasra. Sequential monte carlo samplers. *Journal of the Royal Statistical Society Series B: Statistical Methodology*, 68(3):411–436, 2006.